

# Lecture 2: Minimax theorem, Impagliazzo Hard Core Lemma

Topics in Pseudorandomness and Complexity Theory (Spring 2017)  
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Zero-sum games are two player games played on a matrix  $M \in \text{Mat}_{m \times n}(\mathbb{R})$ . The row player, denoted  $R$ , chooses a row  $i \in [m]$  and the column player  $C$  chooses a column  $j \in [n]$ , simultaneously. The payoff to the row player is  $M_{ij}$  and the payoff to the column player is  $-M_{ij}$  (hence the game is “zero-sum”). We can also consider randomized strategies, where  $R$  chooses a probability distribution  $p$  on  $[m]$ , which can be written as a vector in  $\mathbb{R}^m$ , and the column player chooses a probability distribution  $q$  on  $\mathbb{R}^n$ . The expected payoff to the row player, denoted  $E(p, q)$ , is thus equal to

$$p^T M q = \sum_{ij} M_{ij} p_i q_j.$$

**Theorem 1** (Von Neumann Minimax Theorem). *For a linear game there is a value  $V$  such that*

1. *There exists  $p^*$  such that for all  $q$ ,  $E(p^*, q) \geq V$  and*
2. *there exists a  $q^*$  such that  $E(p, q^*) \leq V$  for all  $p$ .*

In words:  $R$  can guarantee that the expected payoff to  $R$  is at least  $V$ , while  $C$  can guarantee that the expected payoff to  $R$  is at most  $V$ . This  $V$  is called the value of the game.

*Proof.* We say that  $R$  can guarantee that the expected payoff to  $R$  is at least  $\alpha$  if there exists a probability distribution  $p \in \mathbb{R}^m$  such that for all probability distributions  $q \in \mathbb{R}^n$ ,  $E(p, q) \geq \alpha$ .

**Lemma 2.**  *$R$  can guarantee that the expected payoff to  $R$  is at least  $\alpha$  if and only if there is a  $p$  such that for all  $j \in [n]$   $E(p, e_j) \geq \alpha$ .*

*Proof.* The “only” if direction is clear, because the deterministic strategies is a subset of the randomized strategies. To prove the “if” direction, suppose we have such a  $p$ . Note that for any  $q$ ,

$$E(p, q) = \sum_{i \in [m]} \sum_{j \in [n]} p_i M_{ij} q_j = \sum_{j \in [n]} q_j E(p, e_j) \geq \alpha,$$

because  $q$  is a probability distribution on  $[n]$ . □

**Lemma 3.** *Given  $X, Y$  closed, disjoint convex sets in  $\mathbb{R}^n$ , there is a hyperplane that separates  $X$  and  $Y$ , that is, a unit vector  $u$  and a number  $t$  such that either for all  $x \in X$  and  $y \in Y$ ,  $\langle u, x \rangle \leq t$  and  $\langle u, y \rangle > t$  or for all  $x \in X$  and  $y \in Y$ ,  $\langle u, x \rangle < t$  and  $\langle u, y \rangle \geq t$ .*

*Proof sketch.* Suppose  $X$  and  $Y$  are compact. Consider  $\inf_{x \in X, y \in Y} d(x, y)$ ; we know this distance is achieved by some  $x_0$  and  $y_0$  because it is a continuous function on  $X \times Y$  which is also compact. One can check that the hyperplane bisecting the segment  $xy$  separates  $X$  and  $Y$ . Suppose  $X$  and  $Y$  are only closed; then  $X_N \equiv X \cap \{x : |x_i| \leq N\}, Y_N \equiv Y \cap \{x : |x_i| \leq N\}$  are compact sets for each natural number  $N$ . There is a separating hyperplane  $H_N$  for each  $X_N$  and  $Y_N$ ;  $H_N$  can be described by a unit vector  $u_N$  and a threshold  $t_N$  such that  $H_N = \{x : \langle x, u_N \rangle = t_N\}$ . Suppose we always pick  $u_N$  so that  $X_N$  is on the negative side of  $H_N$ .  $t_N$  is in fact the distance from the origin of  $H_N$ , which must be bounded above because some points  $x \in X$  and  $y \in Y$  are at a finite distance from the origin and hence cannot be separated by any hyperplane that is arbitrarily far from the origin despite being contained in  $X_N$  and  $Y_N$  for  $N$  sufficiently large. Hence  $\{(u_N, t_N)\}_{N \in \mathbb{N}}$  resides in a compact subset of  $\mathbb{R}^{n+1}$ , so it has a convergent subsequence. Let  $(u, t)$  be the limit of this subsequence. For each  $x \in X$ , and for  $N$  sufficiently large,  $\langle x, u_N \rangle < t$ , and  $\langle x, u \rangle = \lim_{N \rightarrow \infty} \langle x, u_N \rangle \leq t$ . Similarly for  $y$ , we have  $\langle y, u \rangle \geq t$ . The intersections  $X \cap H$  and  $Y \cap H$  are closed convex sets. If either is empty, we are done. If neither is empty, then by induction there is a hyperplane  $G \subset H$  that separates  $X \cap H$  and  $Y \cap H$  as in the theorem. Without loss of generality suppose  $G$  does not intersect  $X \cap H$ . By translating, assume  $G$  (and so  $H$ ) is through the origin with normal  $v$ , so that  $\langle x, v \rangle < 0$  for  $x \in X$  and  $\langle y, v \rangle \geq 0$  for  $y \in Y$ . Let  $u' = u + .1v$ . Now  $\langle x, u' \rangle = \langle x, u + .1v \rangle < 0$  and similarly  $\langle y, u' \rangle \geq 0$ . Apply the inverse of the translation to obtain a new hyperplane  $(u', t')$  that separates  $X$  and  $Y$  in the way required by the theorem.  $\square$

Take some value  $\alpha$  such that  $R$  cannot guarantee expected value at least  $\alpha$ . Then for all  $p$  there exists  $j$  such that

$$\sum_i p_i M_{ij} \leq \alpha.$$

Let  $X = \{u \in \mathbb{R}^n : u_j \geq \alpha \forall j \in [n]\}$  and  $Y = \{pM : p \text{ is a probability distribution on } [m]\}$ .  $R$  cannot guarantee  $\alpha$  if and only if  $X$  and  $Y$  are disjoint. One needs to check that these are convex and closed. We claim that if  $R$  cannot guarantee  $\alpha$  then there is a hyperplane separating  $X$  and  $Y$ . That is, there is  $l \in \mathbb{R}^n$  and  $b \in \mathbb{R}$  such that for all  $x \in X$ ,  $\langle l, x \rangle > b$  and for all  $y \in T$ ,  $\langle l, y \rangle < b$ . Observe the following:

1. Every entry of  $l$  is nonnegative. Otherwise you could make the corresponding entry in  $x \in X$  arbitrarily large while leaving all others equal to  $\alpha$ . This would violate  $\langle l, x \rangle > b$ .
2. Replace  $l$  from the separating hyperplane by  $q = \frac{l}{\sum l_i}$  so that  $q$  is a probability distribution, and replace  $\alpha$  by  $b = \frac{\alpha}{\sum l_i}$ . So now for all  $x \in X$ ,  $\langle q, x \rangle > b$  and for all  $y \in T$ ,  $\langle q, y \rangle < b$ .
3.  $b \leq \alpha$ , which one can see by applying  $\langle q, x \rangle > b$  for  $x = \alpha \vec{1}$ .

Items 2 and 3 show that for all  $p \in \mathbb{R}^m$ ,  $\langle q, pM \rangle < \alpha$ , or  $E(p, q) < \alpha$ . In other words, if  $C$  plays  $q$  she guarantees that the expected payoff for  $R$  is  $< \alpha$ .

We know that those values  $R$  can guarantee comprise a left half interval, and those values  $C$  can guarantee comprise a right half interval. What we've shown is that for any  $\alpha$  outside  $R$ 's left half interval is in the interior of  $C$ 's right interval. Because our proof was symmetric in  $R$  and  $C$ , any  $\alpha$  outside  $C$ 's right half interval is in the interior of  $R$ 's left half interval. Together these imply the

intervals are closed and intersect in exactly one point.

□

**Theorem 4** (Impagliazzo Hard Core Lemma). *For all  $\epsilon, \delta > 0$ , and  $\Omega(n/\epsilon^2 \log(n/\epsilon^2)) = s \leq 2^{.9n}$ , suppose  $f : \{0, 1\}^n \rightarrow \{0, 1\}$  such that for all circuits  $C$  with size  $s$*

$$\Pr_{x \in \{0,1\}^n} [C(x) = f(x)] < 1 - \delta.$$

*Then there exists a “hard core”  $H \subset \{0, 1\}^n$ ,  $|H| \geq (2\delta)2^n - O(\sqrt{\delta 2^n})$  such that for all circuits  $C$  with size at most  $\epsilon^2 \delta s / 5n$ ,*

$$\Pr_{x \in H} [C(x) = f(x)] < 1/2 + \epsilon.$$

*Proof.* Consider a linear game played by two players  $S$  (sets, row) and  $C$  (circuit, column).  $S$  chooses from sets of  $\{0, 1\}^n$  of size *exactly*  $(2\delta)2^n$ , which we will call “Large sets” and  $C$  chooses from circuits of size at most  $s'$ , which we will call “Small circuits”. The matrix  $M$  of payoffs to the circuit player has entries  $M_{S,C} = \Pr_{x \in S} [C(x) = f(x)]$ .

By the Von Neumann Minimax Theorem and Lemma 2, we are in one of two cases: Either

1. There is a distribution  $\mu_s$  on Large sets such that for all Small circuits  $C$  of size at most  $s'$ ,

$$\mathbb{E}_{S \sim \mu_s} [M_{S,C}] \leq 1/2 + \epsilon,$$

or

2. There is a distribution  $\mu_c$  on Small circuits such that for all Large sets

$$\mathbb{E}_{C \sim \mu_c} [M_{S,C}] \geq 1/2 + \epsilon.$$

### 0.1 Case 1:

We want to obtain a *single* Large set from the distribution  $\mu_s$  on which no circuit agrees well with  $f$ . Define

$$\nu(x) = \Pr_{S \in \mu_s} [x \in S]$$

Then  $\sum_{x \in \{0,1\}^n} \nu(x) = (2\delta)2^n$ , the expected size of  $S$ , by linearity of expectation.

To choose  $H$ , for each  $x \in \{0, 1\}^n$ , add  $x$  to  $H$  independently with probability  $\nu(x)$ . Then

1.  $\Pr[|H| \geq (2\delta)2^n(1 - \eta)] \geq 1 - e^{-\eta^2(2\delta)2^n/3}$ . (The sum of independent Bernoulli's is controlled by the multiplicative Chernoff bound; see Alon and Spencer Corollary A.1.14 or Wikipedia). We will end up choosing  $\eta$  so that this probability is close to one.
2. With high probability over choice of  $H$ ,

$$\Pr_{x \in H} [f(x) = C(x)] \leq 1/2 + \epsilon/2$$

To prove 2, fix a circuit  $C$ . Let  $Y$  be the random variable that is the number of agreements between  $f$  and  $C$  on  $H$ . Then

$$\mathbb{E}_H Y = \sum_{x \in \{0,1\}^n} \Pr[x \in H] 1_{f(x)=C(x)} = \sum_{x \in \{0,1\}^n} \nu(x) 1_{f(x)=C(x)} \quad (1)$$

again by linearity of expectation.  $M_{S,C} = \frac{1}{(2\delta)2^n} \sum_{x \in S} 1_{f(x)=C(x)}$ , so

$$\mathbb{E}_{S \sim \mu_s} [M_{S,C}] = \sum_S \mu_S(S) \frac{1}{(2\delta)2^n} \sum_{x \in S} 1_{f(x)=C(x)} \quad (2)$$

$$= \frac{1}{(2\delta)2^n} \sum_{x \in \{0,1\}^n} 1_{f(x)=C(x)} \sum_{S \ni x} \mu_S(S) \quad (3)$$

$$= \frac{1}{(2\delta)2^n} \sum_{x \in \{0,1\}^n} 1_{f(x)=C(x)} \Pr_{S \sim \mu_s} [x \in S] \quad (4)$$

$$= \frac{1}{(2\delta)2^n} \sum_{x \in \{0,1\}^n} \nu(x) 1_{f(x)=C(x)}. \quad (5)$$

Equations 1 and 5, combined with the fact that we are in Case 1, imply  $\mathbb{E}[Y] = \mathbb{E}_{S \sim \mu_s} [M_{S,C}] (2\delta)2^n \leq (1/2 + \epsilon)(2\delta)2^n$ . Note that  $Y$  is a sum of independent indicator variables, so by the multiplicative Chernoff bound and  $\epsilon < 1/2$

$$\begin{aligned} \Pr[Y > (1/2 + 2\epsilon)(2\delta)2^n] &= \Pr[Y > (1/2 + \epsilon)(2\delta)2^n + \epsilon(2\delta)2^n] \\ &\leq e^{-\frac{1}{3} \left(\frac{\epsilon}{.5+\epsilon}\right)^2 (2\delta)2^n} \leq e^{-\frac{1}{3} \epsilon^2 (2\delta)2^n}. \end{aligned}$$

This was for a fixed  $C$ . By the union bound, with probability at least

$$1 - (\text{number of Small circuits}) e^{-\frac{1}{3} \epsilon^2 (2\delta)2^n} - e^{-\eta^2 (2\delta)2^n / 3} \quad (6)$$

$$= 1 - (s')^{o(s')} e^{-\frac{1}{3} \epsilon^2 (2\delta)2^n} - e^{-\eta^2 (2\delta)2^n / 3} \quad (7)$$

we have that for all Small circuits, the number of agreements between  $C$  and  $f$  on  $H$  is at most  $(1/2 + 2\epsilon)(2\delta)2^n$  and  $|H| \geq (1 - \eta)(2\delta)2^n$ . Provided 7 is positive, there exists such an  $H$ , so we can choose  $s' = \epsilon^2 \delta s / 5n \leq 2^{.9n} \epsilon^2 \delta$  and  $\eta = 5(\delta 2^n)^{-1/2}$  provided  $n$  is large enough.

## 0.2 Case 2:

Now suppose we are in Case 2. We want a small circuit that agrees with  $f$  on a  $(1 - \delta)$  fraction of inputs. First sample  $C_1 \dots C_t$  from  $\mu_C$ . Let  $\eta(x) = \Pr_{C \in \mu_C} [C(x) = f(x)]$ . Recall that  $M_{S,C} = \Pr_{x \in S} [C(x) = f(x)] = \mathbb{E}_{x \in S} [1_{C(x)=f(x)}]$ . So for all Large sets  $S$ ,

$$\begin{aligned} \mathbb{E}_{C \in \mu_C} [M_{S,C}] &= \mathbb{E}_{C \in \mu_C} \mathbb{E}_{x \in S} [1_{C(x)=f(x)}] = \mathbb{E}_{x \in S} \mathbb{E}_{C \in \mu_C} [1_{C(x)=f(x)}] = \\ &= \mathbb{E}_{x \in S} \eta(x) \geq 1/2 + \epsilon. \end{aligned} \quad (8)$$

Let  $n'(x) = \frac{1}{t} \sum_{i=1}^t 1_{C_i(x)=f(x)}$ .

**Claim 5.** With probability at least  $1 - 2^n e^{-.5\epsilon^2 t}$  over choice of  $C_1, \dots, C_t$ , for all  $x$ ,  $|\eta(x) - \eta'(x)| < \epsilon/2$ .

*Proof.* Use Chernoff to bound the probability of the condition not holding for a specific  $x$ : by the additive form of Chernoff,  $\Pr[|\eta'(x) - \eta(x)| \geq \epsilon/2] \leq e^{-.5\epsilon^2 t}$ . Next union bound over  $x \in \{0, 1\}^n$ .  $\square$

If  $t \geq 3n/\epsilon^2$ , there are  $C_1, \dots, C_t$  such that for all  $x$ ,  $|\eta(x) - \eta'(x)| < \epsilon/2$ . By 8,  $\mathbb{E}_{x \in S}[\eta'(x)] \geq 1/2 + \epsilon/2$  for all  $S$  of size  $(2\delta)2^n$ .

A first attempt: Define  $C^*(x) = \text{Maj}(C_1(x), \dots, C_t(x))$ . Let  $S^*$  be the Large set where  $\mathbb{E}_{x \in S}[\eta'(x)]$  is the smallest. In other words,  $S^*$  is the  $(2\delta)2^n$  first  $x$  in increasing order of  $\eta'(x)$ . Note that for  $x \notin S^*$ ,  $\eta'(x) \geq 1/2 + \epsilon/2$ . This implies  $C^*$  is correct outside  $S^*$ , since  $C^*$  is the majority of the  $C_i$  and at least half the  $C_i$  are correct. So  $\Pr_{x \in \{0,1\}^n}[C^*(x) = f(x)] \geq 1 - 2\delta$ , but we need  $1 - \delta$ . It seems like this approach won't tell us much about what happens inside  $S^*$ . To this end, we modify  $C^*$ .

The real attempt: Let  $S^*$  be the defined as in the first attempt. Define  $\beta = \max\{\eta'(x) : x \in S^*\} - 1/2$ . Define

$$C^*(x) = \begin{cases} 0 & \frac{1}{t} \sum C_i(x) \leq \frac{1}{2} - \beta \\ 1 & \frac{1}{t} \sum C_i(x) > \frac{1}{2} + \beta \\ Y & \text{otherwise,} \end{cases} \quad (9)$$

where  $Y$  is chosen independently in  $\{0, 1\}$  with mean  $\frac{1}{2} + \frac{\frac{1}{t} \sum_i C_i(x) - \frac{1}{2}}{2\beta}$ . For  $x \notin S^*$ ,  $\eta'(x) \geq 1/2 + \beta > 1/2$  by the definition of  $\beta$ , so  $C^*$  is correct on  $x$ . Within  $S^*$ , we need to show that the probability over choice of  $x$  and  $Y$  that  $C^*(x)$  is correct is at least  $1/2$ . If  $f(x) = 1$ , then  $\frac{1}{t} \sum C_i(x) = \eta'(x)$  and so

$$\begin{aligned} \Pr_{x \in S^*, Y}[C^*(x) = f(x)] &= \mathbb{E}_{x \in S^*}[\mathbb{E}_Y[C^*(x)]] \\ &\geq E_{x \in S^*} \left[ \frac{1}{2} + \frac{\frac{1}{t} \sum_i C_i(x) - \frac{1}{2}}{2\beta} \right] \\ &= E_{x \in S^*} \left[ \frac{1}{2} + \frac{\eta'(x) - \frac{1}{2}}{2\beta} \right] \geq 1/2 + \epsilon/4\beta > 1/2. \end{aligned}$$

because even if  $\frac{1}{t} \sum_i C_i(x) \leq 1/2 - \beta$ ,  $C^*(x)$  will be zero, which is larger than the negative number  $\frac{1}{2} + \frac{\frac{1}{t} \sum_i C_i(x) - \frac{1}{2}}{2\beta}$ . If  $f(x) = 0$ , then  $\frac{1}{t} \sum C_i(x) = 1 - \eta'(x)$  and so

$$\begin{aligned} \Pr_{x, Y}[C^*(x) = f(x)] &= \mathbb{E}_x[\mathbb{E}_Y[1 - C^*(x)]] \\ &\geq E_{x \in S^*} \left[ \frac{1}{2} + \frac{1/2 - \frac{1}{t} \sum_i C_i(x)}{2\beta} \right] \\ &= E_{x \in S^*} \left[ \frac{1}{2} + \frac{\eta'(x) - \frac{1}{2}}{2\beta} \right] \geq 1/2 + \epsilon/4\beta > 1/2. \end{aligned}$$

If remains to show that  $C^*$  can be computed by a small circuit. This stands to reason, since in addition to the  $t$  copies of  $C$  we just need to add a gadget that outputs 1 if the sum of the  $C_i$  is large enough, 0 if it is small enough, and a random bit with the right probability if it is in between. Intuitively this should take something like  $O(t + ts')$  gates. First, let's show  $C^*$  can be computed by a small randomized circuit  $C(x, r)$  depending on some random bits  $r$ . As  $2\beta$  is  $q/t$  for some integer  $q \in [t]$ , look at

$$\frac{1}{2} + \frac{\frac{1}{t} \sum_i C_i(x) - \frac{1}{2}}{2\beta} \tag{10}$$

$$= \frac{1}{2q} (q + 2 \sum_i C_i(x) - t). \tag{11}$$

This means once we can compute  $q + 2 \sum_i C_i(x) - t$  from the  $C_i$ , we can output 0 if it is  $\leq 0$  and output 1 if it is  $> 2q$ , and if it is in  $[2q]$  output 1 or 0 according to whether it is  $\geq$  or  $<$  a random binary number between 1 and  $2q$  generated by at most  $O(\log t)$  random bits. Since  $q + 2 \sum_i C_i(x) - t$  is at most  $2t$ , we need at most  $O(\log t)$  bits to represent it and  $O(t)$  gates to compute each bit (including a sign bit). To compare the two numbers we only need  $O(\log^2 t)$  gates. In any case, since  $t = 3n/\epsilon^2$ , we need at most

$$3ns'/\epsilon^2 + O(n/\epsilon^2 \log(n/\epsilon^2)) \leq s$$

gates.

If  $C(x, r)$  is a randomized circuit (where  $x$  is the input and  $r$  is random bits) such that  $E_x[E_r[1_{f(x)=C(x,r)}]] \geq 1 - \delta$ , then there exists  $r_0$  such that  $E_x[1_{f(x)=C(x,r_0)}] \geq 1 - \delta$ . Fix that  $r_0$  and hardwire it into  $C$ . This gives a circuit of size at most  $s$  that computes  $f$  with probability at least  $1 - \delta$ .

□